

COMPUTATION OF THE LIMITED INFORMATION
MAXIMUM LIKELIHOOD ESTIMATOR

BY

WARREN DENT

GENE H. GOLUB

STAN-CS-73-339

FEBRUARY 1973

COMPUTER SCIENCE DEPARTMENT
School of Humanities and Sciences
STANFORD UNIVERSITY



COMPUTATION OF THE LIMITED INFORMATION
MAXIMUM LIKELIHOOD ESTIMATOR

by

Warren Dent and Gene H. Golub*

The University of Iowa and Stanford University

*The work of this author was supported in part by the AEC and by NSF.

COMPUTATION OF THE LIMITED INFORMATION MAXIMUM LIKELIHOOD ESTIMATOR

by

Warren Dent and Gene H. Golub

The University of Iowa and Stanford University

0. Abstract

Computation of the Limited Information Maximum Likelihood Estimator (LIMLE) of the set of coefficients in a single equation of a system of interdependent relations is sufficiently complicated to detract from other potentially interesting properties. Although for finite samples the LIMLE has no moments [18], asymptotically it remains normally distributed [2] and retains other properties associated with maximum likelihood. The most extensive application of the estimator has been made in the Brookings studies [7]. We believe that current methods of estimation are clumsy, and present a numerically stable estimation schema based on Householder transformations and the singular value decomposition. The analysis permits a convenient demonstration of equivalence with the Two Stage Least Squares Estimator (TSLS) in the instance of just identification.

1. Introduction

In a system of interdependent relations, suppose a given structural equation is denoted by

$$Y + (X_1, X_2) \begin{pmatrix} \beta \\ \tilde{0} \end{pmatrix} + u = 0.$$

The $n \times (L+1)$ matrix $Y = (Y^*, y)$ represents n observations on L endogenous variables Y^* and n observations on an endogenous variable y elected as subject of the equation. X_1 is the $n \times K_1$ matrix of n observations on the K_1 included exogenous variables, X_2 is the $n \times K_2$ matrix of n observations

on the K_2 excluded exogenous variables, and u is an $n \times 1$ vector of disturbances. $X = (X_1, X_2)$ and is of order $n \times K$, where $K = K_1 + K_2$. γ and β are unknown coefficient vectors, save for the last element of γ which is -1 , whence $\gamma' = (\gamma^*, -1)$.

Assuming X_1 and X of full rank, define the two residual operators

$$M_1 = I - X_1(X_1'X_1)^{-1}X_1',$$

$$M = I - X(X'X)^{-1}X'.$$

The LIMLE is then determined by solving the determinantal equation

$$|Y'M_1Y - \mu Y'MY| = 0$$

for the smallest value of μ , say $\hat{\mu}$. The corresponding solution for $\hat{\gamma}$, with last element -1 , in

$$(Y'M_1Y - \hat{\mu}Y'MY)\hat{\gamma} = 0$$

is the LIMLE of γ , with the LIMLE of β being

$$\hat{\beta} = -(X_1'X_1)^{-1}X_1'Y\hat{\gamma}$$

Derivation of this solution may be found, for example, in [3], [14, pp. 166-173], [6, pp. 335-344], [20, pp. 500-503, pp. 679-686], and [9, pp. 38-44]. The LIMLE is sometimes identified with the Least Variance Ratio Estimator (LVRE), and Least Generalized Residual Variance Estimator (LGRVE), or the Smallest Canonical Correlation Estimator (SCCE). Because of the simplicity of its derivation the LVRE is usually presented in texts, for example [16, pp. 384-387], [15, pp. 166-173], [6, p. 346], [10, p. 338], [17, pp. 567-571], [5, pp. 411-424, pp. 663-666], and [9, pp. 45, 46]. The LGRVE

is proposed in [3], and [10, pp. 338-341], while the SCCE is derived in [4] and [1]. The four estimators are not necessarily identical at all times, since under certain conditions (to be discussed below) on the number of available observations, one may exist while another may not. This confusion has led to some meaningless statements in certain texts, especially with respect to the equivalence of the TSLSE and LIMLE in cases of just-identification.

In the following section we present the computational schema for the LIMLE, while in section 3 we discuss the TSLSE and in section 4 present the determination of the asymptotic variance matrix of these estimators.

2. Algorithm for the LIMLE

There exists an orthogonal $n \times n$ matrix Q , the product of $K \leq n$ Householder transformations [12] such that, for

$$Q = (Q_1, Q_2, Q_3),$$

where Q_1 is $n \times K_1$; Q_2 $n \times K_2$; Q_3 $n \times (n-K)$, Q annihilates X as

$$Q'X = \begin{bmatrix} R_1 & R_3 \\ 0 & R_2 \\ 0 & 0 \end{bmatrix}$$

where R_1 is an upper-triangular non-singular $K_1 \times K_1$ matrix, R_2 is similarly upper-triangular non-singular $K_2 \times K_2$, R_3 is $K_1 \times K_2$, and the zero matrices have appropriate order. Clearly

$$X_1 = Q_1 R_1$$

and

$$X = Q_4 R$$

where

$$Q_4 = (Q_1, Q_2),$$

$$R = \begin{bmatrix} R_1 & R_3 \\ 0 & R_2 \end{bmatrix}.$$

Further, it is readily verified that

$$M = I - Q_4 Q_4'$$

$$= Q_3 Q_3'$$

and

$$M_1 = I - Q_1 Q_1'$$

$$= Q_2 Q_2' + Q_3 Q_3'.$$

Suppose the transformation Q applied to Y yields

$$Q'Y = \begin{bmatrix} Q_1'Y \\ Q_2'Y \\ Q_3'Y \end{bmatrix} = \begin{bmatrix} Z_1 \\ Z_2 \\ Z_3 \end{bmatrix} = (Q'Y^*, Q'y) = \begin{bmatrix} Z_1^*, & z_1 \\ Z_2^*, & z_2 \\ Z_3^*, & z_3 \end{bmatrix}$$

such that matrices with subscript 1 have K_1 rows, while subscript 2 indicates K_2 rows, and subscript 3 indicates $(n-K)$ rows. An asterisk denotes a matrix with L columns, and the partition of Z_1, Z_2, Z_3 corresponds to that of $Y = (Y^*, y)$. Alternately,

$$Y = \sum_{i=1}^3 Q_i Z_i$$

$$Y^* = \sum_{i=1}^3 Q_i Z_i^*,$$

and

$$Y'MY = Z_3'Z_3,$$

$$Y'M_1Y = Z_2'Z_2 + Z_3'Z_3.$$

The determinantal equation for $\hat{\mu}$ now becomes

$$|Z_2'Z_2 + Z_3'Z_3 - \mu Z_3'Z_3| = 0$$

or

$$|Z_2'Z_2 - (\mu-1)Z_3'Z_3| = 0.$$

If the particular equation is over-identified, $K_2 > L$ or $K_2 \geq L+1$ so that Z_2 , which is $K_2 \times (L+1)$, has full column rank with probability one, and the inverse $(Z_2'Z_2)^{-1}$ exists with probability one. We therefore consider

$$|(Z_2'Z_2)^{-1}Z_3'Z_3 - \frac{1}{\mu-1} I| = 0$$

and search for the largest eigenvalue of $(Z_2'Z_2)^{-1}Z_3'Z_3$.

Assuming Z_2 has full rank $L+1$ and since $K_2 \geq L+1$, there exists an orthogonal $K_2 \times K_2$ matrix H , the product of $L+1$ Householder transformations, such that

$$H'Z_2 = \begin{pmatrix} G \\ 0 \end{pmatrix}$$

$$Z_2 = H \begin{pmatrix} G \\ 0 \end{pmatrix} = (H_1, H_2) \begin{pmatrix} G \\ 0 \end{pmatrix} = H_1 G$$

where G is an upper-triangular, non-singular $L+1$ by $L+1$ matrix, and H_1 is $K_2 \times (L+1)$. It follows that

$$Z_2' Z_2 = G' G$$

and

$$(Z_2' Z_2)^{-1} = G^{-1} (G')^{-1}$$

The matrix G^{-1} may be readily computed since G is upper-triangular, see e.g. [19 p. 31, [8, p. 427]. If, now, $n \geq K+L+1$, consider the singular value decomposition [13, 14] of $Z_3 G^{-1}$ as

$$Z_3 G^{-1} = U \Delta V'$$

where U is $(n-K) \times (L+1)$, Δ and V are $(L+1) \times (L+1)$, Δ is diagonal and

$$U' U = I_{L+1} = V' V = V V'.$$

Obviously

$$(G^{-1})' Z_3' Z_3 G^{-1} = V \Delta^2 V'$$

or

$$\begin{aligned} (G^{-1})' Z_3' Z_3 G^{-1} V &= V \Delta^2 \\ &= (\delta_{11}^2 v_1, \delta_{22}^2 v_2, \dots, \delta_{L+1, L+1}^2 v_{L+1}) \end{aligned}$$

where $\delta_1^2 \geq \delta_2^2 \geq \dots \geq \delta_{L+1}^2$, say, are the diagonal elements of Δ^2 , and v_i is the column vector of V corresponding to δ_i^2 . Since

$$G^{-1} (G^{-1})' Z_3' Z_3 G^{-1} v_1 = \delta_1^2 G^{-1} v_1,$$

δ_1^2 is the eigenvalue sought, and we define

$$\hat{\gamma} = G^{-1}v_1.$$

The LIMLE of γ is given by

$$\hat{\gamma} = -\hat{\gamma}/\hat{\gamma}_{L+1}$$

where $\hat{\gamma}_{L+1}$ is the last component of $\hat{\gamma}$.

Since

$$(X_1'X_1)^{-1}X_1' = R_1^{-1}Q_1',$$

$$\begin{aligned}\hat{\beta} &= -(X_1'X_1)^{-1}X_1'Y\hat{\gamma} \\ &= -R_1^{-1}Q_1'(\sum_{i=1}^3 Q_i Z_i)\hat{\gamma} \\ &= -R_1^{-1}Z_1\hat{\gamma}.\end{aligned}$$

When $K \leq n < K+L+1$ the singular value decomposition of $(G^{-1})'Z_3' = U\Delta V'$ is used with

$$\hat{\gamma} = G^{-1}u_1$$

where u_1 is the eigenvector corresponding to the largest value δ_1^2 of δ_i^2 $i=1, \dots, L+1$; and $\hat{\gamma}$ is $\hat{\gamma}$ normalized appropriately. In either case

$$\hat{\mu} = 1+\delta_1^{-2}.$$

3. The LIMLE and TSLSE

When the given structural equation is just-identified, $K_2 = L$. Consider the difference

$$Y'M_1Y - Y'MY = Z_2'Z_2.$$

Since Z_2 has less rows (K_2) than columns ($L+1$), $Z_2'Z_2$ is necessarily singular and

$$|Y'M_1Y - Y'MY| = 0$$

implying $\mu = 1$, which provides the case of the TSLSE, [3], [20, p. 504].

In [5, p. 424] the equivalence of the TSLSE and LIMLE is claimed by assuming $Y'M_1Y = Y'MY$ ($W^*=W$), which obviously need not be true. Indeed there seems to be general confusion as to the behavior of these two residual moment matrices. In [6, p. 339] $Y'M_1Y$ and $Y'MY$ (W_{11} and W_{11}^*) are claimed to be positive definite, while in [15, p. 172] $Y'MY$ ($W_{\Delta\Delta}$) is claimed to be positive definite. Since $Y'MY = Z_3'Z_3$, a necessary condition for non-singularity, with probability one, is that Z_3 have at least as many rows as columns, or that $n \geq K+L+1$. In the derivation so far we have only required $n \geq K$. (This includes derivation of the classical determinantal equation).

Clearly, when $K \leq n < K+L+1$ the LGRVE does not exist (since it is derived through the minimization of a determinant involving $(Y'MY)^{-1}$). Hence equivalence with the TSLSE [10, p. 344] which exists for $n \geq K_1+L+1$ may be impossible in certain instances. Even the LVRE is of spurious interpretation in such instances since the denominator of the ratio being minimized is positive semi-definite and hence may assume a zero value.

For a range of n values ($L+1$ of them) the estimators LVRE, LGRVE, and SCCE may not exist, while the LIMLE will. If the number of included endogenous variables is large for a particular relation in an interdependent system, this will define an equivalent number of observation values over which the LIMLE alone will exist.

4. Asymptotic Covariance

The asymptotic variance-covariance matrix of $(\hat{\gamma}^*, \hat{\beta}')$ is given by the two-stage least squares asymptotic covariance, or by

$$W = s^2 \begin{bmatrix} Y^{*'}(I-M)Y^* & Y^{*'}X_1 \\ X_1'Y^* & X_1'X_1 \end{bmatrix}^{-1}$$

where

$$s^2 = \frac{1}{n} (\underline{y} - Y^*\hat{\gamma}_0^* - X_1\hat{\beta}_0)'(\underline{y} - Y^*\hat{\gamma}_0^* - X_1\hat{\beta}_0)$$

and $\hat{\beta}_0$ is the two-stage least squares estimate of β , while $\hat{\gamma}_0^*$ consists of the first L components of the two-stage least squares estimate of y . The normal equations for the two-stage least squares estimators are

$$\begin{bmatrix} Y^{*'}(I-M)Y^* & Y^{*'}X_1 \\ X_1'Y^* & X_1'X_1 \end{bmatrix} \begin{bmatrix} \gamma_0^* \\ \beta_0 \end{bmatrix} = \begin{bmatrix} Y^{*'}(I-M)\underline{y} \\ X_1'\underline{y} \end{bmatrix}$$

or

$$\begin{bmatrix} Z_1^{*'}Z_1^* + Z_2^{*'}Z_2^* & Z_1^{*'}R_1 \\ R_1'Z_1^* & R_1'R_1 \end{bmatrix} \begin{bmatrix} \gamma_0^* \\ \beta_0 \end{bmatrix} = \begin{bmatrix} Z_1^{*'}z_1 + Z_2^{*'}z_2 \\ R_1'z_1 \end{bmatrix}$$

We now apply Householder transformations as

$$T' \begin{bmatrix} Z_1^* & R_1 & z_1 \\ Z_2^* & 0 & z_2 \end{bmatrix} = \begin{bmatrix} S & t_1 \\ 0 & t_2 \end{bmatrix}$$

reducing the first $L+K_1$ columns to upper triangular form S , whence

$$\begin{bmatrix} \hat{\gamma}_0^* \\ \hat{\beta}_0 \end{bmatrix} = S^{-1} t_1.$$

s^2 may now be computed from above and the asymptotic covariance matrix found as

$$s^2 (S'S)^{-1}.$$

A FORTRAN program listing follows.

Acknowledgement

We wish to thank Mrs. Margaret Wright, at Stanford, for programming the algorithm described in this paper and performing the initial calculations used for checking purposes.

REFERENCES

1. Anderson, T. W. "Estimating Linear Restrictions on Regression Coefficients for Multivariate Normal Distributions," AMS, 1957, Vol. 22, pp. 327-351.
2. Anderson, T. W. "An Asymptotic Expansion of the Distribution of the Maximum Likelihood Estimate of a Coefficient in a Simultaneous Equations System." Presented at the IMS meetings, Montreal, August 1972.
3. Anderson, T. W. and Rubin, H. "Estimation of the Parameters of a Single Equation in a Complete System of Stochastic Equations," AMS, 1949, Vol. 20, pp. 46-63.
4. Bartlett, M. S. "A Note on the Statistical Estimation of Demand and Supply Relationships From Time Series," Econometrica, 1948, Vol. 16, pp. 323-329.
5. Christ, C. F. Econometric Models and Methods, John Wiley & Sons, Inc., New York, 1966.
6. Dhrymes, P. Econometrics, Harper & Row, New York, 1970.
7. Duesenberry, J. S., Fromm, G., Klein, L. R., and Kuh, E. (eds). The Brookings Quarterly Econometric Model of the United States, Rand McNally, Chicago, 1965.
8. Durbin, J. "An Alternative to the Bounds Test for Testing for Serial Correlation in Least-Squares Regression," Econometrica, 1970, Vol. 38, pp. 422-429.
9. Fisk, P. R. Stochastically Dependent Equations, Hafner, New York, 1967.
10. Goldberger, A. S. Econometric Theory, John Wiley & Sons, Inc., New York, 1964.
11. Goldberger, A. S. and Olkin, I. "A Minimum Distance Interpretation of Limited-Information Estimation," Econometrica, 1971, Vol. 39, pp. 635-639.
12. Golub, Gene H. "Numerical Methods for Solving Linear Least-Squares Problems," Numer. Math., 1965, Vol. 7, pp. 206-216.
13. Golub, Gene H. and Kahan, W. "Calculating the Singular Values and Pseudo-inverse of a Matrix," J. Siam Numer. Analys., Ser. B. 1965, Vol. 2. pp. 205-224.
14. Golub, Gene H. and Reinsch, C. "Singular Value Decomposition and Least Squares Solutions." Numer. Math., 1970, Vol. 14, pp. 403-420.

15. Hood, W. C., Koopmans, T. C. (eds). Studies in Econometric Methods, Cowles Foundation Monograph No. 14, John Wiley & Sons, Inc., New York, 1953.
16. Johnston, J. Econometric Methods, McGraw-Hill, New York, 2nd ed., 1972.
17. Kmenta, J. Elements of Econometrics, MacMillan, New York, 1971.
18. Mariano, R. S., and Sawa, T. "Exact Finite-Sample Distribution of Limited-Information Maximum Likelihood Estimator for Two Included Endogenous Variables," JASA, 1972, Vol. 67, pp 159-163.
19. Plackett, R. L. Principles of Regression Analysis, Oxford University Press, 1960.
20. Theil, H. Principles of Econometrics, John Wiley & Sons, Inc., 1971.

```

      IMPLICIT PEAL*8 (A-H,O-Z)
      DIMENSION XY(25,25),S(25),U(25,25),V(25,25)
      PEAL*4 HEAD(20),FMT(20)
      DIMENSION GINV(25,25)
      DIMENSION GAMMA(25),BETA(25),Z3GINV(25,25)
      DIMENSION XXY(625)
      EQUIVALENCE (XY(1,1),XXY(1))
      READ 10,NPROR
      DO 500 NRUN=1,NPROB
      PRINT 726
726  FORMAT(1H1)
c
      READ 7776, HEAD
7776  FORMAT(20A4)
      READ 10,N,K1,K2,L
10    FORMAT(16I5)
c  N IS NUMBER OF ROWS IN ALL MATRICES
c  K1,K2, L ARE THE NUMBERS OF COLUMNS IN X1,X2,Y*, RESPECTIVELY
c
      READ 11,FMT
11    FORMAT(20A4)
c
      KSUM=K1+K2+L+1
c  KSUM IS NUMBER OF COLUMNS IN FULL MATRIX XY
c  THE MATRIX HAS THE FORM
c      (  X1    x2    Y )
c
c      NDIM=25
c  NUMBER OF ROWS DECLARED
      K=K1+K2
      LP1=L+1
      KP1=K+1
      NMK1=N-K1
c
c
c
c  READ THE XY MATRIX FROM CARDS, BY ROWS, IN THE ORDER
c
c      (  X1    x2    Y )
c
      PO 100 I=1,N
      READ (5,FMT) (XY(I,J),J=1,KSUM)
100   CONTINUE
c
c  HEAD IS AN ALPHANUMERIC TITLE TO DESIGNATE THE GIVEN RUN
      PRINT 7775,HEAD
7775  FORMAT(1H1,20A4)
      PRINT 7773, N, K1, K2, L
7773  FORMAT(1H0, 'N=',15, 2X, 'K1=',15,2X, 'K2=',15,2X, 'L=',15)
      PRINT 7772
7772  FORMAT(1H0, 'XY MATRIX')
      DO 602 I=1,N
      PRINT 601,(XY(I,J),J=1,KSUM)
602   CONTINUE

```

```

      IF(K1+L.EQ.K) GO TO 1091
      IF(K1+L.GT.K) GO TO 1092
      PRINT 640
C
C HCOMP1 REDUCES THE FIRST K COLUMNS OF THE RECTANGULAR
C MATRIX XY TO UPPER TRIANGULAR FORM AND APPLIES THE RESULTING
C HOUSEHOLDER TRANSFORMATIONS TO THE REMAINING COLUMNS.
      CALL HCOMP1(NDIM,N,K,KSUM,XY,S)
601  FORMAT(1H0,8D16.6)
640  FORMAT(1H0,/,/)
      DO 667 I=1,K2
      DO 667 J=1,LP1
667  XY(I+K1,J+K1)=XY(I+K1,J+K)
C
C
C NOW COMPUTE HOUSEHOLDER REFLECTIONS TO REDUCE Z2, A K2 BY L+1
C MATRIX WHOSE FIRST LOCATION IS XY(K1+1,K+1), TO UPPER
C TRIANGULAR FORM
      CALL HCOMP1(NDIM,K2,LP1,LP1,XY(K1+1,KP1),S)
C
C CALL THE RESULTING SQUARE TRIANGULAR MATRIX G, WHERE G IS
C (L+1) BY (L+1).
      CALL TNVPT2( LP1,XY(K1+1,KP1),GINV,NDIM,
1      &220)
C
C FORM PRODUCT OF Z3* GINVERSE
      NMK=N-K
      DO 200 I=1,NMK
      DO 200 J=1,LP1
      SUM=0.0D0
      INDEX1 IS INDEX IN XY OF LOCATION PRECEDING FIRST ELEMENT OF Z3
      INDEX1= K*NDIM + K
      DO 300 KK=1,LP1
      IND1= INDEX1 + (KK-1)*NDIM + 1
      SUM = SUM + XXY(IND1)*GINV(KK,J)
C Z3(I, KK)* GINVERSE(KK, J)
300  CONTINUE
      Z3GINV(I,J)=SUM
200  CONTINUE
      NW1=NMK
      NW2=LP1
      IFLAG=0
      IF(N.GE.K+LP1) GO TO 700
      IFLAG=1
      DO 701 I=1,LP1
      II=I+1
      DO 701 J=II,LP1
      WTD=Z3GINV(I,J)
      Z3GINV(I,J)=Z3GINV(J,I)
701  Z3GINV(J,I)=WTD
      IEND=NMK-LP1
      DO 702 I=1,IEND
      NROW=LP1+I
      DO 702 J=1,LP1
702  Z3GINV(NROW,J)=Z3GINV(J,NROW)

```



```

NW1=LP1
NW2=NMK
C
C COMPUTE SINGULAR VALUE DECOMPOSITION OF Z3*GINVERSE
700 CALL DSVD(Z3GINV,NDIM,NDIM,NW1,NW2,0,.TRUE.,.TRUE.,S,U,V)
PRINT 7767
7767 FORMAT(1H0,'MUHAT')
RMU=1.0D0+1.0D0/S(1)**2
PRINT 601, RMU
PRINT 640
RNU=RMU-1.0D0
RNU2=1.0D0/S(2)**2
RMU2=RNU2+1.0D0
IF(IFLAG.EQ.0) GO TO 703
DO 704 I=1,LP1
704 V(I,1)=U(I,1)
C
C SOLVE THE LINEAR SYSTEM G*GAMMA=V, WHERE V IS THE SINGULAR VECTOR
C ASSOCIATED WITH THE LARGEST SINGULARVALUE
C
C NOTE THAT G IS ALREADY UPPER TRIANGULAR
703 CALL TSOLV2(LP1,XY(K1+1,KP1),NDIM,V(1,1),GAMMA,&220)
C
C NORMALIZE GAMMA TO HAVE LAST COMPONENT -1
IF (GAMMA(LP1).EQ.0.0D0) GO TO 220
C G IS (L+1) BY (L+1)
DO 350 I=1,L
GAMMA(I)=-GAMMA(I)/GAMMA(LP1)
350 CONTINUE
C
C COMPUTE RESIDUAL VARIANCE
SQUAR=RMU/(RMU-1.0D0)/GAMMA(LP1)/GAMMA(LP1)/DFLOAT(N)
GAMMA(LP1)=-1.0D0
PRINT 7766
7766 FORMAT(1H0,'GAMMA')
PRINT 601,(GAMMA(I),I=1,LP1)
PRINT 640
C
INDEX1=K*NDIM
C INDEX1 IS THE FIRST LOCATION OF Z1, WHICH IS IN THE FIRST
C ROW, AND THE (K+1)ST COLUMN OF XY
C Z1 IS K1 BY (L+1)
C
C FORM PRODUCT OF THE MATRIX Z1 * GAMMA
DO 360 I=1,K1
SUM=0.0D0
DO 370 KK=1,LP1
IND1=INDEX1+(KK-1)*NDIM+I
C INDEX OF Z1(I, KK)
SUM=SUM+XXY(IND1)*GAMMA(KK)
370 CONTINUE
BETA(I)=SUM
360 CONTINUE
C
C SOLVE R1*BETA = -Z1 * GAMMA

```

```

      CALL TSOLV2(K1,XY,NDIM,BETA,BETA,&220)
      DO 3 870 I=1,K1
      BETA(I)=-BETA(I)
380    CONTINUE
      PRINT 7765
7765  FORMAT(1H0,'BETA')
      PRINT 601,(BETA(I),I=1,K1)
      PRINT 640
      PRINT 707,SQUAR
707  FORMAT(///,' RESIDUAL VARIANCE',D16.6)
C
C COMPUTE ASYMPTOTIC VARIANCES
      NJ=K1+L
      NJ1=NJ+1
      DO 668 I=1,K2
      DO 668 J=1,LP1
668    XY(I+K1,J+K)=XY(I+K1,J+K1)
      DO 677 I=1,LP1
      DO 677 J=1,LP1
      U(I,J)=0.0D0
      DO 677 M=1,NMK1
677    U(I,J)=U(I,J)+XY(M+K1,I+K)*XY(M+K1,J+K)
      DO 669 I=1,K1
      DO 669 J=1,K1
669    GINV(I,J)=XY(I,J)
      DO 670 J=1,L
      DO 670 I=1,K
670    XY(I,J)=XY(I,J+K)
      DO 671 I=1,K1
      DO 671 J=1,K1
671    XY(I,J+L)=GINV(I,J)
      DO 672 J=1,K1
      II=J+1
      DO 672 I=II,K
672    XY(I,J+L)=0.0D0
      DO 678 I=1,K
678    XY(I,NJ1)=XY(I,KSUM)
      CALL HCOMP1(NDIM,K,NJ,NJ1,XY,S)
      CALL TNVRT2(NJ,XY,GINV,NDIM,&220)
      CALL TSOLV2(NJ,XY,NDIM,XY(1,NJ1),S,&220)
      S(LP1)=-1.0D0
      SQUAR=0.0D0
      DO 680 I=1,LP1
      SUM=0.0D0
      DO 679 J=1,LP1
679    SUM=SUM+U(I,J)*S(J)
680    SQUAR=SQUAR+S(I)*SUM
      SQUAR=SQUAR/DFLOAT(N)
      DO 675 I=1,L
      S(I)=0.0D0
      DO 674 J=1,NJ
674    S(I)=S(I)+GINV(I,J)*GINV(I,J)
675    S(I)=S(I)*SQUAR
      DO 676 I=1,K1
      Z3GINV(I,I)=0.0D0

```

```

      II=I+L
      DO 676 J=II,NJ
676   Z3GINV(I,I)=Z3GINV(I,I)+GINV(II,J)*GINV(II,J)
      PRINT 713
713   FORMAT(///' ASYMTOTIC VARIANCES AND Z VALUES FOR GAMMA'/)
      PRINT 601,(S(I),I=1,L)
      DO 714 I=1,L
714   GAMMA(I)=GAMMA(I)/DSQRT(S(I))
      PRINT 601,(GAMMA(I),I=1,L)
      DO 715 I=1,K1
715   Z3GINV(I,I)=Z3GINV(I,I)*SQUAR
      PRINT 716
716   FORMAT(///' ASYMTOTIC VARIANCES AND Z VALUES FOR BETA'/)
      PRINT 601,(Z3GINV(I,I),I=1,K1)
      DO 717 I=1,K1
717   BETA(I)=BETA(I)/DSQRT(Z3GINV(I,I))
      PRINT 601,(BETA(I),I=1,K1)
      KJ=2*(K2-L+1)
      KM=K2-L
      AN=DFLOAT(N)
      TEST1=AN*RNU
      TEST2=AN*DLOG(RMU)
      PRINT 730,TEST1,TEST2,KM
      TEST1=TEST1+AN*RNU2
      TEST2=TEST2+AN*DLOG(RMU2)
      PRINT 731,TEST1,TEST2,KJ
      TEST1=DFLOAT(N-K)*RNU/DFLOAT(K2)
      TEST2=DFLOAT(N-K)*RNU2/DFLOAT(K2)
      PRINT 732,TEST1,TEST2,K2,NMK
730   FORMAT(///// ' N*(MUHAT-1) IS',F12.3,///, ' N*LOG(MUHAT) IS',F11.3,///,
      1' CHI-SQUARE D.F.',I12)
731   FORMAT(///// ' N*(MUHAT(1)+MUHAT(2)-2) IS',F12.3,///, ' N*LOG(MUHAT(1)
      1+MUHAT(2)) IS',F11.3,///, ' CHI-SQUARE D.F.',I12X,I12)
732   FORMAT(///// ' (N-K)*(MUHAT(1)-1)/K2 IS',F12.3,///, ' (N-K)*(MUHAT(2)-
      21)/K2 IS',F12.3,///, ' F DISTRIBUTION D.F.',I12,3X,' ',3X,I4)
      GO TO 500
220   PRINT 221
221   FORMAT(1H0,'SINGULAR UPPER TRIANGULAR MATRIX')
      GO TO 500
1091 PRINT 1093
1093 FORMAT(///' THIS EQUATION IS JUST IDENTIFIED AND TWO-STAGE'/' LEAST
      1 SQUARES IS APPROPRIATE')
      GO TO 500
1092 PRINT 1094
1094 FORMAT(///' THIS EQUATION IS NOT IDENTIFIABLE')
500   CONTINUE
      STOP
C
C
C
      END

```

```

SUBROUTINE DSVD(A,MMAX,NMAX,M,N,P,WITHU,WITHV,S,U,V)
IMPLICIT REAL*8 (A-H,O-Z)
DIMENSION A(MMAX,NMAX),U(MMAX,NMAX),V(NMAX,NMAX)
DIMENSION S(N),B(100),C(100),T(100)

```

THIS SUBROUTINE COMPUTES THE SINGULAR VALUE DECOMPOSITION OF A REAL M*N MATRIX A. I. E. IT COMPUTES MATRICES U, S, AND V SUCH THAT

$$A = U * S * V^T,$$

WHERE

U IS AN M*N MATRIX AND $U^T * U = I$, (U^T =TRANPOSE OF U),

V IS AN N*N MATRIX AND $V^T * V = I$, (V^T =TRANPOSE OF V),

AND S IS AN N*N DIAGONAL MATRIX,

DESCRIPTION OF PARAMETERS:

A = REAL*8 ARRAY. A CONTAINS THE MATRIX TO BE DECOMPOSED.

MMAX = INTEGER*4 VARIABLE. THE NUMBER OF DECLARED ROWS IN THE ARRAYS A AND U.

NMAX = INTEGER*4 VARIABLE. THE NUMBER OF DECLARED ROWS IN THE ARRAY V.

M, N = INTEGER*4 VARIABLES. THE NUMBER OF ROWS AND COLUMNS IN THE MATRIX STORED IN A. ($N \leq M \leq 100$. IF IT IS NECESSARY TO SOLVE A LARGER PROBLEM, THEN THE AMOUNT OF STORAGE ALLOCATED TO THE ARRAYS B, C, AND T MUST BE INCREASED ACCORDINGLY.)

INTEGER P

LOGICAL WITHU, WITHV

WITHU, WITHV = LOGICAL*4 VARIABLES. IF WITHU=.TRUE., THEN THE MATRIX U IS COMPUTED AND STORED IN THE ARRAY U. SIMILARLY FOR V.

S = REAL*8 ARRAY. S(1), . . . , S(N) CONTAIN THE DIAGONAL ELEMENTS OF THE MATRIX S ORDERED SO THAT $S(1) \geq S(1+1)$, $1=1, \dots, N-1$.

U, V = REAL*8 ARRAYS. U, V CONTAIN THE MATRICES U AND V. IF WITHU=.TRUE. AND WITHV=.FALSE., THEN THE ACTUAL PARAMETER CORRESPONDING TO A AND U MAY BE THE SAME. SIMILARLY FOR V IF WITHV=.TRUE. AND WITHU=.FALSE..

P = INTEGER*4 VARIABLE. IF $P > 0$, THEN COLUMNS $N+1, \dots, N+P$ OF A ARE ASSUMED TO CONTAIN THE COLUMNS OF AN M*P MATRIX B. THIS MATRIX IS MULTIPLIED BY U^T , AND UPON EXIT, A CONTAINS IN THESE SAME COLUMNS THE N*P MATRIX $U^T * B$. ($P \geq 0$)

THIS SUBROUTINE IS A TRANSLATION OF AN ALGOL 60 PROCEDURE DESCRIBED IN THE ARTICLE "SINGULAR VALUE DECOMPOSITION AND

C LEAST SQUARES SOLUTIONS, NUM. MATH. 14 (1970), PP. 403-420.
 C THE TRANSLATION WAS DONE BY P. BUSINGER AT BELL TELEPHONE
 C LABORATORIES WITH SOME CHANGES AND EDITING DONE BY R.
 C UNDERWOOD AT STANFORD UNIVERSITY.
 C

DATA ETA /Z3410000000000000/
 DATA TOL /Z0D100000000000000/

C ETA AND TOL ARE MACHINE DEPENDENT CONSTANTS WHOSE
 C VALUES ARE $16^{**(-13)}$ AND $16^{**(-52)}$, RESPECTIVELY,
 C ON IBM SYSTEM/360 COMPUTERS.
 C

NP=N+P
 N1=N+1

C HOUSEHOLDER REDUCTION TO BIDIAGONAL FORM

C(1)=0.0D0

K=1

10 K1=K+1

C ELIMINATION OF A(I,K), I=K+1, . . . , M

Z=0.0D0

DO 20 I=K,M

20 Z=Z+A(I,K)**2

B(K)=0.0D0

IF (Z.LE.TOL) GOTO 70

Z=DSORT(Z)

B(K)=Z

W=DABS(A(K,K))

Q=1.0D0

IF (W.NE.0.0D0) Q=A(K,K)/W

A(K,K)=Q*(Z+W)

IF (K.EQ.NP) GOTO 70

DO 50 J=K1,NP

Q=0.0D0

DO 30 I=K,M

30 Q=Q+A(I,K)*A(I,J)

Q=Q/(Z*(Z+W))

DO 40 I=K,M

40 A(I,J)=A(I,J)-Q*A(I,K)

50 CONTINUE

C PHASE TRANSFORMATION

Q=-A(K,K)/DABS(A(K,K))

DO 60 J=K1,NP

60 A(K,J)=Q*A(K,J)

C ELIMINATION OF A(K,J), J=K+2, ..., N

70 IF (K.EQ.N) GOTO 140

Z=0.0D0

DO 80 J=K1,N

80 Z=Z+A(K,J)**2

C(K1)=0.0D0

IF (Z.LE.TOL) GOTO 130

```

      Z=DSORT(Z)
      C(K1)=Z
      W=DABS(A(K,K1))
      Q=1.0D0
      IF (W.NE.0.0D0) Q=A(K,K1)/W
      A(K,K1)=Q*(Z+W)
      DO 110 I=K1,M
        Q=0.0D0
        PO 90 J=K1,N
      90   Q=Q+A(K,J)*A(I,J)
        Q=Q/(Z*(Z+W))
        DO 100 J=K1,N
      100   A(I,J)=A(I,J)-Q*A(K,J)
      110   CONTINUE
C
C   PHASE TRANSFORMATION
      Q=-A(K,K1)/DABS(A(K,K1))
      DO 120 I=K1,M
      120   A(I,K1)=A(I,K1)*Q
C
      130 K=K1
      GOTO 10
C
C   TOLERANCE FOR NEGLIGIBLE ELEMENTS
      140 EPS=0.0D0
      DO 150 K=1,N
        S(K)=R(K)
        T(K)=C(K)
      150   EPS=DMAX1(EPS,S(K)+T(K))
      EPS=EPS*ETA
C
C   INITIALIZATION OF U AND V
      IF (.NOT.WITHU) GOTO 180
      DO 170 J=1,N
        PO 160 I=1,M
      160   U(I,J)=0.0D0
      170   U(J,J)=1.0D0
C
      180 IF (.NOT.WITHV) GOTO 210
      PO 200 J=1,N
        DO 190 I=1,N
      190   V(I,J)=0.0D0
      200   V(J,J)=1.0D0
C
C   QR DIAGONALIZATION
      210 DO 380 KK=1,N
        K=N1-KK
C
C   TEST FOR SPLIT
      220 DO 230 LL=1,K
        L=K+1-LL
        IF (DABS(T(L)).LE.EPS) GOTO 290
        IF (DABS(S(L-1)).LE.EPS) GOTO 240
      230   CONTINUE

```

C CANCELLATION

```
240 CS=0.0D0
    SN=1.0D0
    L1=L-1
    DO 280 I=L,K
        F=SN*T(I)
        T(I)=CS*T(I)
        IF (DABS(F).LE.EPS) GOTO 290
        H=S(I)
        W=DSORT(F*F+H*H)
        S(I)=W
        CS=H/W
        SN=-F/W
        IF (.NOT.WITHU) GOTO 260
    DO 250 J=1,N
        X=U(J,L1)
        Y=U(J,I)
        U(J,L1)=X*CS+Y*SN
250   U(J,I)=Y*CS-X*SN
260   IF (NP.EQ.N) GOTO 280
    DO 270 J=N1,NP
        Q=A(L1,J)
        R=A(I,J)
        A(L1,J)=Q*CS+R*SN
270   A(I,J)=P*CS-Q*SN
280 CONTINUE
```

C
C TEST FOR CONVERGENCE
290 W=S(K)
 IF (L.EQ.K) GOTO 360

C
C ORIGIN SHIFT
 X=S(L)
 Y=S(K-1)
 G=T(K-1)
 H=T(K)
 F=((Y-W)*(Y+W)+(G-H)*(G+H))/(2.0D0*H*Y)
 G=DSORT(F*F+1.0D0)
 IF (F.LT.0.0D0) G=-G
 F=((X-W)*(X+W)+(Y/(F+G)-H)*H)/X

C
C OR STEP
 CS=1.0D0
 SN=1.0D0
 L1=L+1
 PO 350 I=L1,K
 G=T(I)
 Y=S(I)
 H=SN*G
 G=CS*G
 W=DSORT(H*H+F*F)
 T(I-1)=W
 CS=F/W
 SN=H/W
 F=X*CS+G*SN

```

      G=G*CS-X*SN
      U=Y*SN
      Y=Y*CS
      IF (.NOT.WITHV) GOTO 310
      DO 300 J=1,N
        X=V(J,I-1)
        W=V(J,I)
        V(J,I-1)=X*CS+W*SN
300      V(J,I)=W*CS-X*SN
310      W=DSORT(H*H+F*F)
      S(I-1)=W
      CS=F/W
      SN=H/W
      F=CS*G+SN*Y
      X=CS*Y-SN*G
      IF (.NOT.WITHU) GOTO 330
      PO 320 J=1,N
        Y=U(J,I-1)
        W=U(J,I)
        U(J,I-1)=Y*CS+W*SN
320      U(J,I)=W*CS-Y*SN
330      IF (N.EQ.NP) GOTO 350
      PO 340 J=N1,NP
        Q=A(I-1,J)
        R=A(I,J)
        A(I-1,J)=Q*CS+R*SN
340      A(I,J)=R*CS-Q*SN
350      CONTINUE

C
      T(L)=0.0D0
      T(K)=F
      S(K)=X
      GOTO 220

C
C
C      CONVERGENCE
360      IF (W.GE.0.0D0) GOTO 380
      S(K)=-W
      IF (.NOT.WITHV) GOTO 380
      DO 370 J=1,N
370      V(J,K)=-V(J,K)
380      CONTINUE

C
C
C      SORT SINGULAR VALUES
      DO 450 K=1,N
        G=-1.0D0
        J=K
        DO 390 I=K,N
          IF (S(I).LE.G) GOTO 390
          G=S(I)
          J=I
390      CONTINUE
          IF (J.EQ.K) GOTO 450
          S(J)=S(K)
          S(K)=G
          IF (.NOT.WITHV) GOTO 410
          PO 400 I=1,N

```



```

      Q=V(I,J)
      V(I,J)=V(I,K)
400    V(I,K)=Q
410    IF (.NOT.WITHU) GOTO 430
      PO 420 I =1,N
      Q=U(I,J)
      U(I,J)=U(I,K)
420    U(I,K)=Q
430    IF (N.EQ.NP) GOTO 450
      no 440 I=N1,NP
      Q=A(J,I)
      A(J,I)=A(K,I)
440    A(K,I)=Q
450    CONTINUE

```

```

C
c    BACK TRANSFORMATION
      IF (.NOT.WITHU) GOTO 510
      DO 500 KK=1,N
      K=N1-KK
      IF (B(K).EQ.0.0D0) GOTO 500
      Q=-A(K,K)/DABS(A(K,K))
      DO 460 J=1,N
460    U(K,J)=Q*U(K,J)
      DO 490 J=1,N
      Q=0.0D0
      PO 470 I=K,M
470    Q=Q+A(I,K)*U(I,J)
      Q=Q/(DABS(A(K,K))*B(K))
      PO 480 I=K,M
      U(I,J)=U(I,J)-Q*A(I,K)
480
490    CONTINUE
500    CONTINUE

```

```

C
510  IF (.NOT.WITHV) GOTO 570
      IF (N.LT.2) GOTO 570
      PO 560 KK=2,N
      K=N1-KK
      K1=K+1
      IF (C(K1).EQ.0.0D0) GOTO 560
      Q=-A(K,K1)/DABS(A(K,K1))
      PO 520 J=1,N
520    V(K1,J)=Q*V(K1,J)
      PO 550 J=1,N
      Q=0.0D0
      DO 530 I=K1,N
530    Q=Q+A(K,I)*V(I,J)
      Q=Q/(DABS(A(K,K1))*C(K1))
      DO 540 I=K1,N
      V(I,J)=V(I,J)-Q*A(K,I)
540
550    CONTINUE
560    CONTINUE

```

```

c
570  RETURN
      END

```

SUBROUTINE TNVRT2 (N, A, AI, IDIM, *)

INTEGER N, IDIM
REAL*8 A(IDIM,N), AI(IDIM,N)

C TNVRT2 IS A MODIFICATION OF INVRT2 TO INVERT AN
C ORIGINAL MATRIX THAT IS UPPER TRIANGULAR.
C CALLS TO DECOMP2 AND IMPRV2 ARE OMITTED.

EXTERNAL TSOLV2
INTEGER I, J
REAL*8 E(100), X(100)

DO 3 I=1,N
E(I) = 0.0D0
3 CONTINUE
DO 1 I=1,N
E(I) = 1.0D0
CALL TSOLV2 (N, A, IDIM, E, X, &10)
DO 2 J=1,N
AI(J,I) = X(J)
2 CONTINUE
E(I) = 0.0D0
1 CONTINUE
RETURN
10 RETURN 1

C LAST CARD OF SUBROUTINE TNVRT2.
END

SUBROUTINE TSOLV2 (N, LU, IDIM, B, X, *)

INTEGER N, IDIM

REAL*8 LU(IDIM,N), B(N), X(N)

C SOLVES $LU \cdot X = B$, WHERE LU IS UPPER TRIANGULAR.
C TSOLV2 IS A MODIFICATION OF THE USUAL SOLVE2, WHICH FOLLOWS
C DECOMP2.

C
C

INTEGER I, J, I P, I P1, IM1, NP1, I RACK

REAL*8 SUM

NP1 = N + 1

IF (LU(N,N).EQ.0.000) RETURN 1

X(N) = B(N)/LU(N,N)

IF(N.EQ.1) RETURN

C BACK SUBSTITUTION

DO 4 I RACK =2,N

I = NP1 - IBACK

C I GOES FROM (N-1) TO 1 .

I P1 = I+1

SUM = 0.000

DO 3 J=I P1,N

SUM = SUM + LU(I,J)*X(J)

3 CONTINUE

IF (LU(I,I).EQ.0.000) RETURN 1

X(I) = (B(I)-SUM)/LU(I,I)

4 CONTINUE

RETURN

C LAST CARD OF SUBROUTINE TSOLV2.
END

```

SUBROUTINE HCOMP1(MDIM,M,N,NTOTAL,A,U)
C
  INTEGER MDIM,M,N
  DOUBLE PRECISION A(MDIM,N),U(M)
C
C  HOUSEHOLDER REDUCTION OF THE FIRST N COLUMNS OF A
C  TO UPPER TRIANGULAR FORM.
C  HCOMP1 IS TAKEN FROM CLEVE MOLEP'S SUBROUTINE HECOMP,
C  WHICH REDUCES ALL COLUMNS OF THE MATRIX TO TRIANGULAR FORM.
C  HCOMP1 WAS DESIGNED FOR A SPECIAL PURPOSE, TO MAKE THE FIRST
C  N COLUMNS UPPER TRIANGULAR, WHILE APPLYING THE TRANSFORMATIONS
C  TO NTOTAL COLUMNS.
C
C
C  MDIM= DECLARED ROW DIMENSION OF A
C  M= NUMBER OF ROWS
C  N= NUMBER OF COLUMNS OF A WHOSE UPPER PORTIONS ARE TO
C  BE REDUCED TO TRIANGULAR FORM. AFTER REDUCTION, THE
C  MATRIX A WILL CONTAIN AN N BY N UPPER TRIANGULAR MATRIX
C  IN ITS UPPER LEFT CORNER.
C  NTOTAL= TOTAL NUMBER OF COLUMNS OF A. THE TRANSFORMATIONS
C  WILL BE APPLIED TO ALL COLUMNS.
C  FOR STANDARD LEAST-SQUARES PROBLEMS, NTOTAL=N.
C
C  A= M BY NTOTAL MATRIX WITH M.GE.NTOTAL
C  INPUT, MATRIX TO BE REDUCED
C  OUTPUT, REDUCED MATRIX AND INFORMATION ABOUT REDUCTION
C  U= VECTOR OF LENGTH M
C  INPUT, IGNORED
C  OUTPUT, INFORMATION ABOUT REDUCTION
C
C
C  DOUBLE PRECISION ALPHA,BETA,GAMMA,DSORT
C
C  P O G K=1,N
C
C  FIND REFLECTION THAT ZEROS A(I,K),I=K+1,...,M
C
C  ALPHA=0.0
C  DO 1 I=K,M
C    U(I)=A(I,K)
C    ALPHA=ALPHA+U(I)**2
1  CONTINUE

```

```

ALPHA=DSQRT(ALPHA)
IF(U(K).LT.0.0) ALPHA=-ALPHA
U(K)=U(K)+ALPHA
BETA=ALPHA*U(K)
A(K,K)=-ALPHA
IF(BETA.EQ.0.0 .OR. K.EQ.NTOTAL) GO TO 6

```

C
C
C

APPLY REFLECTION TO REMAINING COLUMNS OF A

```

      KP1=K+1
      DO 4 J=KP1,NTOTAL
        GAMMA=0.0
        DO 2 I=K,M
          GAMMA=GAMMA+U(I)*A(I,J)
2        CONTINUE
        GAMMA=GAMMA / BETA
        DO 3 I=K, M
          A(I,J)=A(I,J) - GAMMA * U(I)
3        CONTINUE
4      CONTINUE
6    CONTINUE
      RETURN

```

C
C
C
C

TRIANGULAR RESULT IS STORED IN A(I,J),U(I),E,J
VECTORS DEFINING REFLECTIONS ARE STORED IN U AND REST OF A

END