

ON A MODEL FOR COMPUTING ROUND-OFF ERROR OF A SUM

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Given real numbers  $a_1, a_2, \dots, a_n$  we are interested in the classic problem of the error in computing  $S = \sum_{i=1}^n a_i$  when the sum is computed by  $\tilde{s}_0 = \sum_{i=1}^n a_i^*$  where  $a_i^*$  is the nearest integer to  $a_i$ . We shall first study this error as a function of a A shift, i.e., when all numbers  $a_i$  are each shifted A and then rounded;

$$(1) \quad S - nA = \sum_{i=1}^n (a_i - A)$$

$$(2) \quad \tilde{s}_\Delta - nA = \sum_{i=1}^n (a_i - A)^*$$

We will then let A become a random variable that can take on uniformly any value in the interval  $-\frac{1}{2} \leq A \leq +\frac{1}{2}$ . Different choices of A give rise to different rounding errors  $\tilde{s}_\Delta - s$  and the variance of the distribution of  $\tilde{s}_\Delta - s$  can be used to measure the variability of the rounding error due to the random selection of the origin of the real numbers  $a_i$  with respect to that of the computer.

The cumulative error from (1) and (2) is

$$(3) \quad \tilde{s}_\Delta - s = \sum_{i=1}^n [(a_i - \Delta)^* - (a_i - A)]$$

Let  $f_i$  be the positive fractional part of  $a_i$  and let  $\alpha_i$  be the largest integer not exceeding  $a_i$ , i.e.,

$$(4) \quad a_i = \alpha_i + f_i$$

Denoting by  $r_i$  the error of the  $i^{\text{th}}$  term, we have

$$(5) \quad r_i = [(a_i - A)^* - (a_i - A)] = \begin{cases} 1 - (f_i - \Delta) & \text{if } -\frac{1}{2} \leq A \leq -\frac{1}{2} + f_i \\ -(f_i - \Delta) & \text{if } -\frac{1}{2} + f_i \leq A \leq +\frac{1}{2} \end{cases}$$

To prove the above, we note that  $f_i - \Delta = (a_i - \Delta) + \alpha_i$ . If  $-\frac{1}{2} \leq f_i - \Delta \leq +\frac{1}{2}$  then  $(a_i - \Delta)$  is rounded to  $a_i$ . Hence  $a_i - \Delta$  is rounded down if  $-\frac{1}{2} + f_i \leq A$  otherwise rounded up.

Denoting expected value by  $E$ , we have by direct evaluation

$$(6) \quad E(r_i) = \int_{-\frac{1}{2}}^{+\frac{1}{2}} r_i d\Delta = 0$$

Assume  $f_i \leq f_j$ , then

$$\begin{aligned} E(r_i r_j) &= \int_{-\frac{1}{2}+f_i}^{-\frac{1}{2}+f_j} r_i r_j dA + \int_{-\frac{1}{2}+f_i}^{-\frac{1}{2}+f_j} r_i r_j d\Delta + \int_{-\frac{1}{2}+f_j}^{+\frac{1}{2}} r_i r_j d\Delta \\ &= \int_{-\frac{1}{2}}^{+\frac{1}{2}} (f_i f_j - \Delta(f_i + f_j) + \Delta^2) d\Delta \\ &+ \int_{-\frac{1}{2}}^{-\frac{1}{2}+f_i} [(1-f_i-f_j) + 2\Delta] d\Delta \\ &+ \int_{-\frac{1}{2}+f_j}^{-\frac{1}{2}+f_i} [-f_i + A] d\Delta \end{aligned}$$

Performing indicated integration yields:

$$(7) \quad E(r_i r_j) = \frac{1}{2} [ |f_j - f_i|^2 - |f_j - f_i| + \frac{1}{6} ]$$

which is one-half the 2<sup>nd</sup> order Bernoulli Polynomial in  $|f_j - f_i|$ . For  $f_j < f_i$  we also get (7). Note that the individual errors  $r_i$  and  $r_j$  are not independent of one another.

It now follows that

$$(9) \quad E(S) = S$$

$$(10) \quad E(\tilde{S}-S)^2 = E\left(\sum_{i=1}^n \sum_{j=1}^n r_i r_j\right) = \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n [|f_i - f_j|^2 - |f_i - f_j| + \frac{1}{6}]$$

The usual value of variance,  $E(S-S)^2 = n/12$ , will result if we further assume  $f_i$  are independently drawn from uniform distributions on  $[0 \leq f_i \leq 1]$ .

Theorem: If the fractional parts of all  $a_i$  are equal to each other, then each term of (10) is maximum for  $0 \leq f_i \leq 1$  and

$$(11) \quad \text{Max } E(S-S)^2 = \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n \left(\frac{1}{6}\right) = \frac{n^2}{12} .$$

From (10) we have an interesting inequality, namely for all  $f_i$

$$(12) \quad V(f) = \frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n \{|f_i - f_j|^2 - |f_i - f_j| + \frac{1}{6}\} \geq 0$$

This function is not convex even for  $n=2$ , since  $f^{(1)} = (\frac{1}{2}, 0)$  and  $f^{(2)} = (-\frac{1}{2}, 0)$  yields  $V(f^0) = V(f^1) = \frac{1}{12} + \frac{1}{12} - \frac{1}{12} = \frac{1}{12}$  but

$V\left(\frac{f^1 + f^2}{2}\right) = V(0) = \frac{3}{12}$ . There appears to be no obvious direct way to establish that  $V(f) \geq 0$  for all  $0 \leq f_i \leq 1$ . Our development shows  $V(f)$  to be a variance and this, of course, constitutes an indirect proof. We can replace (12) by a convex realization: Assume  $f_i \geq f_{i+1}$  for all  $i$ , then the problem of finding  $\text{Min } V(f)$  can be rewritten:

$$(13) \text{ Find } \text{Min } [V(f)] = \sum_{i < j} (f_i - f_j)^2 + \frac{n^2}{12} - [(n-1)f_1 + (n-3)f_2 + (n-5)f_3 + \dots + (n-2k+1)f_k + \dots - (n-1)f_n]$$

subject to

$$(14) \quad f_1 \geq f_2 \geq \dots \geq f_n$$

$$(15) \quad 0 \leq f_i \leq 1$$

Formally (13), (14), (15), is a positive definite quadratic program. Fortunately, as we shall see this can be solved by classical calculus by ignoring inequalities (14) and (15).

Theorem: Equally spaced  $f_i = (n - i)/n$ , ( $i = 1, \dots, n$ ) yields  $\text{Min } V(f) = \frac{1}{12}$  independent of  $n$ , i.e., the variance of the sum in this case is minimum and is the same as the variance of the individual terms forming the sum.

Proof: Setting partials = 0 in (13) yields:

$$(16) \left\{ \begin{array}{l} 2(n-1)f_1 - 2f_2 \dots - 2f_n = (n-1) \\ - 2f_1 + 2(n-1)f_2 \dots - 2f_n = (n-3) \\ - 2f_1 - 2f_2 \dots 2(n-1)f_{n-1} - 2f_n = -(n-3) \\ - 2f_1 - 2f_2 \dots 2(n-1)f_n = -(n-1) \end{array} \right.$$

Adding shows the equations to be dependent. Hence we may drop the last equation as redundant. Moreover, we can always translate the  $f_1$  so that the smallest  $f_i$ , namely  $f_n = 0$

Re-adding yields:

$$2f_1 + 2f_2 + \dots + 2f_{n-1} + 0 = (n-1) \quad , f_n = 0.$$

Adding this last equation to each of the others gives

$$2nf_1 = (n - 2i + 1) + (n - 1) = 2(n - i)$$

$$(17) \quad f_i = (n - i)/n$$

Evidently the conditions  $0 \leq f_i \leq 1$  and  $f_i \geq f_{i+1}$  are (by good luck) also satisfied so that (17) yields the minimum, namely

$$(18) \quad \text{Min } V(f) = \frac{n^2}{12} - \frac{1}{2} \sum_{i=1}^n (n-2i+1)f_i = \frac{1}{12} ,$$